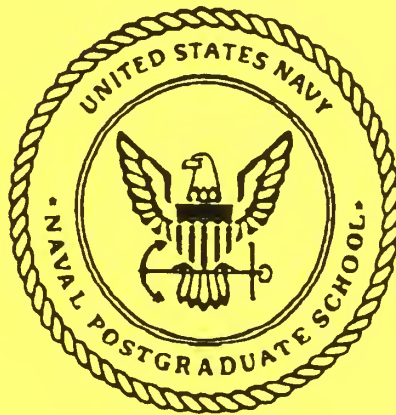


523

# NAVAL POSTGRADUATE SCHOOL

## Monterey, California



A COMPARISON OF PREDICTORS FOR  
FIRST-GUESS WIND SPEED ERRORS

Donald P. Gaver  
Patricia A. Jacobs

December 1993

Approved for public release; distribution is unlimited.

Prepared for:  
Naval Research Laboratory-West  
Monterey, CA 93943-5006

FedDocs  
D 208.14/2  
NPS-OR-93-020



NAVAL POSTGRADUATE SCHOOL  
MONTEREY, CA 93943-5000

Rear Admiral T. A. Mercer  
Superintendent

Harrison Shull  
Provost

This report was prepared for and funded by Naval Research Laboratory-West.

Reproduction of all or part of this report is authorized.

This report was prepared by:

# REPORT DOCUMENTATION PAGE

Form Approved  
OMB No. 0704-0188

Public reporting burden for this collection of information is estimated to average 1 hour per response, including the time for reviewing instructions, searching existing data sources, gathering and maintaining the data needed, and completing and reviewing the collection of information. Send comments regarding this burden estimate or any other aspect of this collection of information, including suggestions for reducing this burden, to Washington Headquarters Services, Directorate for Information Operations and Reports, 1215 Jefferson Davis Highway, Suite 1204, Arlington, VA 22202-4302, and to the Office of Management and Budget, Paperwork Reduction Project (0704-0188), Washington, DC 20503.

1. AGENCY USE ONLY (Leave blank)	2. REPORT DATE December 1993	3. REPORT TYPE AND DATES COVERED Technical Report
----------------------------------	---------------------------------	--

4. TITLE AND SUBTITLE A Comparison of predictors for first-guess wind speed errors	5. FUNDING NUMBERS RH2G5
---	-----------------------------

6. AUTHOR(S) Donald P. Gaver and Patricia A. Jacobs
--

7. PERFORMING ORGANIZATION NAME(S) AND ADDRESS(ES) Naval Postgraduate School Monterey, CA 93943	8. PERFORMING ORGANIZATION REPORT NUMBER NPS-OR-93-020
---	---

9. SPONSORING / MONITORING AGENCY NAME(S) AND ADDRESS(ES) Naval Research Laboratory-West Monterey, CA 93943-5006	10. SPONSORING / MONITORING AGENCY REPORT NUMBER
--	--

11. SUPPLEMENTARY NOTES
-------------------------

12a. DISTRIBUTION / AVAILABILITY STATEMENT Approved for public release; distribution is unlimited.	12b. DISTRIBUTION CODE
---	------------------------

13. ABSTRACT (Maximum 200 words) Numerical meteorological models are used to assist in the prediction of weather. Each run of a numerical model produces forecasts of meteorological variables which are used as preliminary predictions of the future values of these variables. These initial predictions are referred to as <i>first-guess</i> values. Estimation of the mean-square first-guess error is required in the optimal interpolation process in the numerical prediction of atmospheric variables. Several predictors for the mean-square error of the first-guess wind speeds are studied. The results suggest that prediction using observed covariates tend to be better than those using first-guess covariates. However, observed covariates are not always available. Predictions using first-guess covariates are better at the 250 mb level than the 850 or 500 mb levels. Of those first-guess covariates studied, first-guess wind speed appears to be the best.
---

14. SUBJECT TERMS Gaussian model with log-linear scale parameter; nonparametric models; prediction of mean square errors; first-guess errors in meteorological models; generalized linear regression	15. NUMBER OF PAGES 36
	16. PRICE CODE

17. SECURITY CLASSIFICATION OF REPORT Unclassified	18. SECURITY CLASSIFICATION OF THIS PAGE Unclassified	19. SECURITY CLASSIFICATION OF ABSTRACT Unclassified	20. LIMITATION OF ABSTRACT UL
---	--	---	----------------------------------

# A Comparison of Predictors For First-Guess Wind Speed Errors

by P. A. Jacobs and D. P. Gaver

## Abstract

Numerical meteorological models are used to assist in the prediction of weather. Each run of a numerical model produces forecasts of meteorological variables which are used as preliminary predictions of the future values of these variables. These initial predictions are referred to as *first-guess* values. Estimation of the mean-square first-guess error is required in the optimal interpolation process in the numerical prediction of atmospheric variables. Several predictors for the mean-square error of the first-guess wind speeds are studied. The results suggest that prediction using observed covariates tend to be better than those using first-guess covariates. However, observed covariates are not always available. Predictions using first-guess covariates are better at the 250 mb level than the 850 or 500 mb levels. Of those first-guess covariates studied, first-guess wind speed appears to be the best.

## 1. INTRODUCTION AND SUMMARY

Numerical meteorological models are used to assist in the prediction of weather. Each run of a numerical model produces forecasts of meteorological variables which are used as preliminary predictions of future values of these variables. These initial predictions are referred to as first-guess values. In this paper first-guess values will refer to the most recent 12-hour forecasts.

In certain areas of the world, observations of forecasted variables become available. Prior to the next run of the numerical model a multivariate optimal interpolation analysis updates a first-guess value of a variable by

adding to it a weighted observed value of the variable if it is available. The weight multiplying the observed value depends on estimates of the mean-squared error of the first-guess value and the mean-squared error of the observation; cf. Goerss et al., [1991, a, b]. Thus it is of importance to predict the first-guess mean-squared errors.

The general problem of modeling and predicting mean-square errors is important but not widely studied; see Davidian and Carroll (1987), Nelder and Lee (1992), Aitken (1987), McCullagh and Nelder (1983).

In Jacobs and Gaver (1991, 1992) statistical models for the error of the first guess are used to predict mean-square error for first-guess wind components. The models assume that the error of the first guess has a normal distribution with mean 0 and variance which is a function that is log-linear with covariates. Details of the model are presented in Appendix A.

In this paper we use data from February 1991 to compare the predictive ability of various models. The data consist of measurements and 12 hour forecasts (first-guess values) of  $u$  and  $v$  wind components at the 850 mb, 500 mb and 250 mb pressure levels from 93 stations in North America 25N–75N for the month of February 1991. The forecasts are produced using the NOGAPS Spectral Forecast Model; cf. Hogan et al., (1991). Each station has measurement and first-guess values for every 12 hours; there are some missing observations. These missing values are deleted from the data set. The measurement values are subtracted from the first-guess values to obtain observations of the error of the first-guess value.

Let  $U(o;t)$ , (respectively  $V(o;t)$ ), be the observed  $u$ -wind, (respectively  $v$ -wind) component at time  $t$ . Let  $U(f;t)$ , (respectively  $V(f;t)$ ), be the first-guess  $u$ -wind (respectively  $v$ -wind) component at time  $t$ ;  $U(f;t)$ , (respectively  $V(f;t)$ )

is the forecasted value of the  $u$ -wind (respectively  $v$ -wind) component made 12 hours previously. The first-guess error for the  $u$ -wind (respectively  $v$ -wind) component is

$$Y_U(t) = U(f;t) - U(o;t); \text{ (respectively } Y_V(t) = V(f;t) - V(o;t)). \quad (1.1)$$

The following covariates are considered in the log-linear model for the mean-square error of the first guess.

$$r(o;t) = \left[ (U(o;t) - U(o;t-1))^2 + (V(o;t) - V(o;t-1))^2 \right]^{\frac{1}{2}} \quad (1.2)$$

$$w(o;t) = \left[ U(o;t)^2 + V(o;t)^2 \right]^{\frac{1}{2}} \quad (1.3)$$

$$r(f;t) = \left[ (U(f;t) - U(f;t-1))^2 + (V(f;t) - V(f;t-1))^2 \right]^{\frac{1}{2}} \quad (1.4)$$

$$w(f;t) = \left[ U(f;t)^2 + V(f;t)^2 \right]^{\frac{1}{2}}. \quad (1.5)$$

$$r^*(t) = \left[ (U(f;t) - U(o;t-1))^2 + (V(f;t) - V(o;t-1))^2 \right]^{\frac{1}{2}} \quad (1.6)$$

$$a(o,U;t) = |U(o;t) - U(o;t-1)|, \quad a(o,V;t) = |V(o;t) - V(o,t-1)| \quad (1.7)$$

$$a(f,U;t) = |U(f;t) - U(f;t-1)|, \quad a(f,V;t) = |V(f;t) - V(f,t-1)| \quad (1.8)$$

$$a^*(U;t) = |U(f;t) - U(o;t-1)|, \quad a^*(V;t) = |V(f;t) - V(o,t-1)| \quad (1.9)$$

$$m(f;t) = \max(U(f;t), V(f;t)) \quad (1.10)$$

The resultant wind  $r(o;t)$ , (respectively  $r(f;t)$  and  $r^*(t)$ ), is a measure of the observed (respectively forecasted), change in the wind. The variable  $w(o;t)$ , (respectively  $w(f;t)$ ), is the observed, (respectively forecasted), wind speed. Higher wind speeds suggest more activity in the atmosphere. The change in magnitudes  $a(o,U;t)$ ,  $a(f,U;t)$  and  $a^*(U;t)$  (respectively  $a(o,V;t)$ ,  $a(f,V;t)$  and  $a^*(V;t)$ ) will be used to predict  $Y_U(t)$ , (respectively  $Y_V(t)$ ).

The data are randomly divided into two sets called DA and DB. Maximum likelihood estimates of the parameters of the models using different covariates are computed using data DA (respectively DB). Nonparametric models based on binning are also considered. The models are then used to predict the mean-square first-guess errors in data set DB (respectively DA). Log-likelihood functions and the empirical distribution of the first-guess errors normalized by their predicted mean-square errors are used to evaluate the models' predictive ability. Details are given in Section 2.

In general, models which use observed covariates, e.g.  $w(o)$ ,  $a(o)$ , have more predictive ability than those that use first-guess covariates, e.g.  $w(f)$ ,  $a(f)$ ,  $m(f)$ . The models applied at the 250 mb level appear to have more predictive ability than those for 500 mb and 850 mb.

Among the one-variate models for the 250 mb pressure height, the models that statistically appear to have the most predictive ability have as their covariate  $w(o)$ ,  $a(o)$  or  $r(o)$ . Those that have less but some predictive ability have as their covariate  $a^*$ ,  $w(f)$ ,  $m(f)$  or  $r^*$ . Finally, one-variate models using variates  $r(f)$  and  $a(f)$  appear to have little predictive ability. Among those models for the 250 mb pressure height that use one first-guess covariate,  $m(f)$  or  $w(f)$  appear to have the most predictive ability.



## 2. THE DATA ANALYSIS

In this section we describe the data analysis. Let  $U_i(o;t)$  and  $U_i(f;t)$ , (respectively  $V_i(o;t)$  and  $V_i(f;t)$ ) be the observed and first-guess  $u$ -wind (respectively  $v$ -wind) component at location  $i = 1, \dots, S$  at time  $t$ . By data we mean the vector  $(U_i(o,t), U_i(f,t), V_i(o,t), V_i(f,t), U_i(o,t-1), U_i(f,t-1), V_i(o,t-1), V_i(f,t-1))$ . The data set contains missing values. Vectors containing these missing values are deleted from the data set. Once missing values are deleted, there are 3618 vectors at the 850 mb level, 4100 at the 500 mb level, and 3744 at the 250 mb level. The observed values are subtracted from the first-guess values to obtain observations of the first-guess errors for each wind component

$$Y_i(U;t) = U_i(f;t) - U_i(o;t)$$

$$Y_i(V;t) = V_i(f;t) - V_i(o;t).$$

The remaining data are randomly divided into two sets called DA and DB without regard to the values of the data, the time  $t$ , or the location. Thus, data from the same location for different times may be in different data sets. Models are estimated for each pressure level using only covariates for that pressure level. The covariates considered for each wind component appear in Appendix B. The general statistical model is described in Appendix A.

The model is estimated using data sets DA, DB, and all the data for each pressure level. The estimated values for the parameters for selected models appear in Tables 3A, 4A, 3B, 4B, 3C and 4C. Note that the parameter estimates are usually positive. Hence increased values of the covariates are associated with higher variance of the first-guess errors.

The models estimated from DA (respectively DB) are used to predict the first-guess errors in data set DB (respectively DA). One measure used for

assessing a model's goodness of fit and predictive ability is the value of  $\tilde{\ell}$ , the log-likelihood function up to addition of constants given in Appendix A (A.4); the log-likelihood for predicting mean-square errors in DB using a model estimated using DA uses the first-guess error and covariate(s) from DB and the parameter estimates from DA. Values of  $\tilde{\ell}$  are computed for data DA (respectively DB) using the parameters estimated using DB (respectively DA); these values assess each model's predictive ability. Values of  $\tilde{\ell}$  are also computed for data DA (respectively DB) using parameters estimated using DA (respectively DB); these values assess each model's goodness of fit.

Tables 1A, 1B, 1C present the values of  $\tilde{\ell}$  for one-variate models for the different pressure levels. Also displayed are the values of  $\tilde{\ell}$  for a model in which the first-guess errors are independent normally distributed with mean 0 and constant variance  $e^\alpha$ .

Tables 2A, 2B, 2C present values for  $\tilde{\ell}$  for two-variate models.

Compare the value of  $\tilde{\ell}_c$  for the model with constant variance (no covariates) for DA (respectively DB) fit using DA (respectively DB) with the values of  $\tilde{\ell}$  for DA (respectively DB) using models with parameters estimated using the other half of the data DB (respectively DA). A value of  $\tilde{\ell}$  greater than  $\tilde{\ell}_c$  indicates that the corresponding model fit with the other half of the data describes the data better than the best constant variance model fit with the same data it is used to summarize. For 850 mb data those one-variate models for which  $\tilde{\ell} > \tilde{\ell}_c$  for DA and DB for both wind components are those with variate  $r^*, a^*, r$ , and  $w(o)$ . For 500 mb the one-variate models are those with variate  $r^*, a(o), r, w(o)$ , and  $w(f)$ . For 250 mb, the models are those with variate  $r^*, a(o), r, w(o), w(f)$ , and  $m(f)$ .

To compare the predictive ability of the models, the fraction of increase in  $\bar{\ell}$ ,  $(\bar{\ell} - \bar{\ell}_c)/|\bar{\ell}_c|$  is computed where  $\bar{\ell}_c$  is the maximum value of  $\bar{\ell}$  for the constant variance model (with no covariates) estimated using data DA (respectively DB) compared to the value of  $\bar{\ell}$  for DA using one-variate models estimated using the other half of the data DB (respectively DA). The values of percentage of increase appear in Table 5 for the one-variable models with variate  $r^*$ ,  $a^*$ ,  $r$ ,  $w(o)$ ,  $w(f)$ , and  $m(f)$ . Note that the fraction increase tends to be larger for the 250 mb for the covariates using observed data,  $r$ ,  $w(o)$ , and  $a(o)$ . The fraction also tends to be larger for the first-guess covariates  $w(f)$  and  $m(f)$  at the 250 mb level.

Another measure of predictability is the distribution of the first-guess errors divided by their predicted standard deviations. Table 6 displays the moments of the first-guess errors of the wind components DA (respectively DB) divided by the standard deviations that are predicted for it using the model fit using data of DB (respectively DA). Recall that the models assume that these errors are normally distributed with mean 0. Thus, if a model were perfect then the mean (respectively standard deviation, skewness, and kurtosis) of the normalized first-guess errors would be 0, (respectively 1, 0 and 3). Of particular interest is the kurtosis. In this application, the kurtosis can be thought of as a measure of the variability of the variance (cf. Cramér page 356). Hence, the smaller the kurtosis, the better the prediction of the model.

Table 6 presents not only results for the model of Appendix A with various covariates but also results for a nonparametric one-variate model. This nonparametric model is as follows. The data in DA (respectively DB) are binned into  $N$  bins according to the value of the ordered covariate. For each bin, the mean of the square of the wind speed errors corresponding to the

covariates in that bin is computed. To evaluate the predictive ability of the model, the other data set DB (respectively DA) is used. The predicted mean-square error for a data point in DB (respectively DA) is the mean of the square of the wind speed errors for the bin determined from DA (respectively DB) the data point's covariate lies in.

Table 6 presents selected results using data for the 250 mb level. Results for parametric models of Appendix A with parameters estimated by maximum likelihood (MLE) and the nonparametric models with bins are presented. Also displayed are the sample moments of the first-guess errors in the row labeled "none". Displayed in the row labeled "constant" are the sample moments for the first-guess errors divided by the predicted standard deviation for a model with constant variance  $e^\alpha$  fit using the other half of the data.

The values of the kurtosis suggest the following. Once again, models using the observed covariates  $a(o)$  and  $r(o)$  appear to make the best predictors. Among the first-guess covariates,  $m(f)$  and  $w(f)$  appear to have comparable predictive ability. If a nonparametric model using a first-guess variate is being considered, then using first-guess wind speed as the covariate seems to be a good choice.

## REFERENCES

- Aitken, M., (1987). "Modeling variance heterogeneity in normal regression using GLIM." *Appl Stat.*, **36**, pp. 332-339.
- Cox, D. R. and D. V. Hinkley, *Theoretical Statistics*, Chapman and Hall, London, 1974.
- Cramér, H., *Mathematical Methods of Statistics*, Princeton University Press, Princeton, NJ, 1946.
- Davidian, M., and R. J. Carroll (1987). "Variance function estimation," *J. Am Statist. Ass.*, **82**, pp. 1079-1091.
- de Bruijn, N. G., *Asymptotic Methods in Analysis*, Interscience, New York, 1958.
- Easton, G. S., "Location compromise maximum likelihood estimators," in *Configural Polysampling*, ed. S. Morgenthaler and J. W. Tukey, Wiley, New York, 1991, pp. 157-192.
- Goerss, J. S. and P. A. Phoebus (1991a). *The Multivariate Optimum Interpolation Analysis of Meteorological Data at FNOC*. NOARL Report Number 31, Naval Oceanographic and Atmospheric Research Laboratory, Stennis Space Center, MS.
- Goerss, J. S. and P. A. Phoebus (1991b). "The Navy's operational atmospheric analysis." to appear in *Weather and Forecasting*.
- Hogan, T. F., and T. E. Rosmond (1991). "The description of the Navy operational global atmospheric prediction system's spectral forecast model." *Monthly Weather Review*, **119**, No. 8., pp. 1786-1815.
- Jacobs, P. A. and D. P. Gaver, "Preliminary results from the analysis of wind component error." *Naval Postgraduate School Technical Report NPSOR-91-029*, September, 1991.
- Jacobs, P. A. and D. P. Gaver, "Preliminary results from the analysis of wind component error—July Data." *Naval Postgraduate School Technical Report*, to appear.
- McCullagh and J. A. Nelder, *Generalized Linear Regression*, Chapman and Hall, New York, 1983.
- Nelder, J. A. and Y. Lee (1992), "Likelihood, quasi-likelihood and pseudo-likelihood: some comparisons." *J. R. Statist. Soc., B*, **54**, No. 1, pp. 273-284.

APPENDIX A  
THE STATISTICAL MODEL

In this Appendix we describe the statistical model. Let  $U_i(o;t)$  (respectively  $V_i(f;t)$ ) denote the observed  $u$ -wind component (respectively first-guess wind component) at location  $i$  at time  $t$ ;  $i = 1, \dots, S$ . Let  $V_i(o;t)$ , (respectively  $V_i(f;t)$ ) denote the observed  $v$ -wind component (respectively first-guess wind component) at location  $i$  at time  $t$ . The first-guess error of the  $u$ -wind (respectively  $v$ -wind) component at location  $i$  at time  $t$  is

$$Y_i(U;t) = U_i(o;t) - U_i(f;t)$$

(respectively,

$$Y_i(V;t) = V_i(o;t) - V_i(f;t).$$

(A.1)

The model is that  $\{Y_i(U;t), i = 1, \dots, S\}$  and  $\{Y_i(V;t), i = 1, \dots, S\}$  are independent random variables having a normal distribution with mean 0. The variance of  $Y_i(U;t)$  is log-linear with a number of covariates. That is

$$\begin{aligned} & \text{Var}[Y_i(U;t) | X_i(1;t) = x_i(1), \dots, X_i(p;t) = x_i(p)] \\ & = \exp\left\{\alpha + \sum_{j=1}^p \beta_j(t)x_i(j;t)\right\}. \end{aligned} \quad (\text{A.2})$$

The likelihood function for this model is (up to multiplication by constants)

$$L(\alpha, \beta_1, \dots, \beta_p)$$

$$= \prod_t \prod_i \exp\left\{-\frac{1}{2}\left[\alpha + \sum_{j=1}^p \beta_j(t)x_i(j;t)\right]\right\} \exp\left\{-\frac{1}{2}y^2 \exp\left\{-\left[\alpha + \sum_{j=1}^p \beta_j(t)x_i(j;t)\right]\right\}\right\} \quad (\text{A.3})$$

The log-likelihood function is (up to addition by constants)

$$\begin{aligned} & \bar{\ell}(\alpha, \beta_1, \dots, \beta_p) \\ &= \sum_t \sum_i -\frac{1}{2} \left[ \alpha + \sum_{j=1}^p \beta_j(t) x_i(j;t) \right] - \frac{1}{2} y^2 \exp \left\{ - \left( \alpha + \sum_{j=1}^p \beta_j(t) x_i(j;t) \right) \right\} \end{aligned} \quad (\text{A.4})$$

The recursive procedure used to estimate the parameters  $(\alpha, \beta_1, \dots, \beta_p)$  is described in Gaver and Jacobs [1991].

APPENDIX B  
THE COVARIATES

In this Appendix we list the covariates that were considered. As before let  $U_i(o;t)$  and  $U_i(f;t)$ , (respectively  $V_i(o;t)$  and  $V_i(f;t)$ ) denote the observed and first-guess  $u$ -wind (respectively  $v$ -wind) component at time  $t$  for location  $i$ ,  $i = 1, \dots, S$ . The covariates considered for the first-guess error of the  $u$ -wind component are

$$a_i(o,U;t) = |U_i(o;t) - U_i(o;t-1)|$$

$$a_i(f,U;t) = |U_i(f;t) - U_i(f;t-1)|$$

$$a_i^*(U;t) = |U_i(f;t) - U_i(o;t-1)|$$

$$w_i(o;t) = [U_i(o;t)^2 + V_i(o;t)^2]$$

$$w_i(f;t) = [U_i(f;t)^2 + V_i(f;t)^2]$$

$$r_i(o;t) = \left[ [U_i(o;t) - U_i(o;t-1)]^2 + [V_i(o;t) - V_i(f;t-1)]^2 \right]^{\frac{1}{2}}$$

$$r_i(f;t) = \left[ [U_i(f;t) - U_i(f;t-1)]^2 + [V_i(o;t) - V_i(f;t-1)]^2 \right]^{\frac{1}{2}}$$

$$r_i^*(t) = \left[ [U_i(f;t) - U_i(o;t-1)]^2 + [V_i(f;t) - V_i(o;t-1)]^2 \right]^{\frac{1}{2}}$$

$$m(f;t) = \max(U(f;t), V(f;t)).$$

The covariates considered for the first-guess error of the  $v$ -wind component are



$$a_i(o, V; t) = |V_i(o; t) - V_i(o; t - 1)|$$

$$a_i(f, V; t) = |V_i(f; t) - V_i(f; t - 1)|$$

$$a_i^*(V; t) = |V_i(f; t) - V_i(o; t - 1)|$$

$w_i(o; t), w_i(f; t), r_i(o; t), r_i(f; t),$  and  $r_i^*(t)$ .

TABLE 1A  
Log-Likelihood  
850 mb. Height/February Data

One-Variate Models

Wind Comp	Data Model	Const	$r^*(t)$	$r(f;t)$	$a(o;t)$	$a(f;t)$	$a^*(t)$	$r(t)$	$w(o;t)$	$w(f;t)$	$m(f;t)$
u	A	-6226.6	-6183.6	-6226.6	-6109.6	-6226.6	-6167.1	-6162.3	-6174.4	-6217.1	-6221.5
	B	-6439.8	-6331.3	-6434.9	-6327.7	-6403.8	-6307.6	-6377.8	-6312.1	-6412.9	-6439.0
	B	-6452.9	-6347.4	-6452.5	-6341.0	-6450.1	-6319.6	-6390.2	-6333.9	-6426.7	-6455.0
	A	-6238.8	-6196.0	-6241.6	-6121.9	-6257.0	-6176.1	-6173.8	-6193.8	-6228.6	-6235.7
v	A	-6445.9	-6356.9	-6441.4	-6440.5	-6443.6	-6362.4	-6406.9	-6275.3	-6435.8	-6442.9
	B	-6343.7	-6297.4	-6343.1	-6287.4	-6343.5	-6317.6	-6258.1	-6201.8	-6327.9	-6343.7
	B	-6346.5	-6302.2	-6347.3	-6316.4	-6349.4	-6326.4	-6271.6	-6207.2	-6332.2	-6336.7
	A	-6448.8	-6362.4	-6446.0	-6470.7	-6450.5	-6373.7	-6422.0	-6280.7	-6440.4	-6451.4

$$r(f;t) = \left[ (U(f;t) - U(f;t-1))^2 + (V(f;t) - V(f;t-1))^2 \right]^{1/2}$$

$$a(f;t) = |U(f;t) - U(f;t-1)|$$

$$r^*(t) = \left[ (U(f;t) - U(o;t-1))^2 + (V(f;t) - V(o;t-1))^2 \right]$$

$$a(o;t) = |U(o;t) - U(o;t-1)| \quad \text{for } u \text{ - wind component error}$$

$$a(o;t) = |V(o;t) - V(o;t-1)| \quad \text{for } v \text{ - wind component error}$$

$$a^*(t) = |U(f;t) - U(o;t-1)| \quad \text{for } u \text{ - wind component error}$$

$$a^*(t) = |V(f;t) - V(o;t-1)| \quad \text{for } v \text{ - wind component error}$$

$$r(t) = \left[ (U(o;t) - U(o;t-1))^2 + (V(o;t) - V(o;t-1))^2 \right]^{1/2}$$

$$w(o;t) = \left[ U(o;t)^2 + V(o;t)^2 \right]^{1/2}$$

$$w(f;t) = \left[ U(f;t)^2 + V(f;t)^2 \right]^{1/2}$$

$$m(f;t) = \max(U(f;t), V(f;t))$$

TABLE 2A  
 Log-Likelihood  
 850 mb. Height/February Data

Two-Variate Models

Wind Comp	Data Model	Const	$r^*(t), a(o;t)$	$r(f;t), a(f;t)$	$r^*(t), a^*(t)$	$r^*(t), a(f;t)$	$r(t), w(o;t)$	$r(t), w(f;t)$
u	A	-6226.6	-6085.7	-6226.6	-6164.6	-6177.0	-6137.0	-6160.4
	B	-6439.8	-6240.0	-6394.5	-6299.0	-6329.4	-6294.8	-6362.8
	B	-6452.9	-6262.5	-6449.6	-6312.4	-6367.9	-6318.9	-6379.3
	A	-6238.8	-6102.3	-6263.4	-6174.8	-6202.6	-6158.7	-6173.6
v	A	-6445.9	-6356.8	-6441.1	-6353.6	-6344.2	-6260.1	-6404.2
	B	-6343.7	-6271.4	-6339.1	-6297.1	-6267.8	-6181.7	-6257.1
	B	-6346.5	-6304.1	-6345.6	-6306.5	-6277.0	-6190.3	-6270.8
	A	-6448.8	-6396.3	-6448.7	-6365.0	-6355.1	-6270.2	-6419.7

$$r(f;t) = \left[ (U(f;t) - U(f;t-1))^2 + (V(f;t) - V(f;t-1))^2 \right]^{1/2}$$

$$a(f;t) = |U(f;t) - U(f;t-1)|$$

$$r^*(t) = \left[ (U(f;t) - U(o;t-1))^2 + (V(f;t) - V(o;t-1))^2 \right]$$

$$a(o;t) = |U(o;t) - U(o;t-1)| \quad \text{for } u \text{ - wind component error}$$

$$a(o;t) = |V(o;t) - V(o;t-1)| \quad \text{for } v \text{ - wind component error}$$

$$a^*(t) = |U(f;t) - U(o;t-1)| \quad \text{for } u \text{ - wind component error}$$

$$a^*(t) = |V(f;t) - V(o;t-1)| \quad \text{for } v \text{ - wind component error}$$

$$r(t) = \left[ (U(o;t) - U(o;t-1))^2 + (V(o;t) - V(o;t-1))^2 \right]^{1/2}$$

$$w(o;t) = \left[ U(o;t)^2 + V(o;t)^2 \right]^{1/2}$$

$$w(f;t) = \left[ U(f;t)^2 + V(f;t)^2 \right]^{1/2}$$

TABLE 3A  
 One-Variate Models  
 Parameter Estimates  
 (Standard Errors)  
 850 mb. February Data

Wind Comp	Data Set	$r^*(t)$		$r(f;t)$		$a(o;t)$		$a(f;t)$		$a^*(t)$	
		$\alpha$	$\beta$	$\alpha$	$\beta$	$\alpha$	$\beta$	$\alpha$	$\beta$	$\alpha$	$\beta$
u	A	2.08 (0.06)	0.04 (0.007)	2.44 (0.06)	0.0005 (0.007)	2.00 (0.05)	0.08 (0.008)	2.44 (0.05)	0.002 (0.01)	2.14 (0.05)	0.07 (0.009)
	B	2.00 (0.06)	0.07 (0.007)	2.46 (0.06)	0.02 (0.007)	2.15 (0.05)	0.08 (0.008)	2.33 (0.05)	0.06 (0.01)	2.10 (0.05)	0.09 (0.009)
	All	2.03 (0.04)	0.06 (0.005)	2.45 (0.04)	0.009 (0.005)	2.08 (0.04)	0.08 (0.006)	2.37 (0.03)	0.04 (0.008)	2.11 (0.04)	0.08 (0.006)
v	A	2.07 (0.06)	0.06 (0.007)	2.47 (0.05)	0.02 (0.007)	2.48 (0.05)	0.02 (0.008)	2.51 (0.05)	0.01 (0.008)	2.20 (0.05)	0.06 (0.007)
	B	2.13 (0.06)	0.05 (0.007)	2.47 (0.06)	0.006 (0.008)	2.22 (0.05)	0.06 (0.008)	2.52 (0.05)	-0.004 (0.008)	2.29 (0.05)	0.04 (0.008)
	All	2.10 (0.04)	0.05 (0.005)	2.47 (0.04)	0.01 (0.005)	2.35 (0.03)	0.04 (0.005)	2.51 (0.03)	0.005 (0.006)	2.24 (0.04)	0.05 (0.005)

$$r(f;t) = \left[ (u(f;t) - u(f;t-1))^2 + (v(f;t) - v(f;t-1))^2 \right]^{1/2}$$

$$a(f;t) = |u(f;t) - u(f;t-1)|$$

$$r^*(t) = \left[ (u(f;t) - u(o;t-1))^2 + (v(f;t) - v(o;t-1))^2 \right]$$

$$a(o;t) = |u(o;t) - u(o;t-1)|, \quad a^*(t) = |u(f;t) - u(o;t-1)|$$

TABLE 4A

Two-Variate Models  
Parameter Estimates  
(Standard Errors)

850 mb. February Data

Wind Comp	Data Set	log MSE = $\alpha + \beta_1 r^*(t) + \beta_2 a(o;t)$		log MSE = $\alpha + \beta_1 r_f(t) + \beta_2 a(f;t)$		log MSE = $\alpha + \beta_1 r^*(t) + \beta_2 a^*(t)$				
		$\alpha$	$\beta_2$	$\alpha$	$\beta_2$	$\alpha$	$\beta_2$			
u	A	1.75 (0.07)	0.04 (0.007)	0.08 (0.008)	2.13 (0.06)	0.05 (0.008)	-0.03 (0.012)	2.08 (0.06)	0.02 (0.01)	0.05 (0.01)
	B	1.65 (0.07)	0.06 (0.007)	0.07 (0.008)	1.98 (0.06)	0.06 (0.008)	0.02 (0.01)	1.99 (0.06)	0.03 (0.009)	0.07 (0.01)
	All	1.70 (0.05)	0.05 (0.005)	0.08 (0.006)	2.04 (0.05)	0.06 (0.006)	-0.004 (0.008)	2.03 (0.04)	0.02 (0.007)	0.06 (0.009)
v	A	2.08 (0.07)	0.06 (0.007)	-0.002 (0.009)	2.47 (0.06)	0.02 (0.01)	-0.009 (0.02)	2.09 (0.06)	0.04 (0.01)	0.03 (0.01)
	B	2.04 (0.07)	0.03 (0.008)	0.04 (0.009)	2.45 (0.06)	0.03 (0.02)	-0.03 (0.02)	2.13 (0.06)	0.05 (0.01)	-0.007 (0.01)
	All	2.05 (0.05)	0.05 (0.005)	0.02 (0.006)	2.46 (0.04)	0.03 (0.01)	-0.02 (0.01)	2.10 (0.04)	0.05 (0.009)	0.01 (0.01)

$$r(f;t) = \left[ (u(f;t) - u(f;t-1))^2 + (v(f;t) - v(f;t-1))^2 \right]^{1/2}$$

$$a(f;t) = |u(f;t) - u(f;t-1)|$$

$$r^*(t) = \left[ (u(o;t) - u(o;t-1))^2 + (v(f;t) - v(o;t-1))^2 \right]$$

$$a^*(t) = |u(o;t) - u(o;t-1)|, \quad a^*(t) = |u(f;t) - u(o;t-1)|$$

TABLE 1B  
Log-Likelihood  
500 mb. Height/February Data

One-Variate Models

Wind Comp	Data Model	Const	$r^*(t)$	$r(f;t)$	$a(o;t)$	$a(f;t)$	$a^*(t)$	$r(t)$	$w(o;t)$	$w(f;t)$	$m(f;t)$
u	A	-7894.4	-7865.0	-7877.6	-7699.3	-7876.1	-7866.0	-7732.5	-7846.3	-7866.9	-7860.5
	B	-7965.1	-7816.5	-7928.9	-7711.9	-7963.0	-7854.9	-7780.2	-7890.7	-7936.8	-7923.2
	B	-7966.4	-7854.2	-7933.2	-7714.0	-7973.1	-7871.6	-7780.7	-7893.5	-7937.9	-7925.4
	A	-7895.6	-7897.9	-7881.8	-7701.0	-7885.5	-7878.3	-7732.9	-7848.9	-7868.0	-7862.9
v	A	-7720.5	-7702.4	-7716.8	-7624.9	-7709.3	-7695.9	-7616.2	-7683.2	-7696.0	-7706.7
	B	-7849.7	-7841.9	-7847.6	-7723.5	-7845.6	-7842.2	-7740.4	-7780.6	-7827.2	-7828.2
	B	-7853.8	-7849.2	-7852.2	-7727.3	-7853.3	-7854.3	-7744.2	-7787.3	-7830.9	-7832.9
	A	-7724.5	-7709.1	-7721.2	-7628.3	-7715.9	-7706.3	-7619.8	-7689.3	-7699.5	-7711.2

$$r(f;t) = \left[ (U(f;t) - U(f;t-1))^2 + (V(f;t) - V(f;t-1))^2 \right]^{1/2}$$

$$a(f;t) = |U(f;t) - U(f;t-1)|$$

$$r^*(t) = \left[ (U(f;t) - U(o;t-1))^2 + (V(f;t) - V(o;t-1))^2 \right]$$

$$a(o;t) = |U(o;t) - U(o;t-1)| \quad \text{for } u \text{ - wind component error}$$

$$a(o;t) = |V(o;t) - V(o;t-1)| \quad \text{for } v \text{ - wind component error}$$

$$a^*(t) = |U(f;t) - U(o;t-1)| \quad \text{for } u \text{ - wind component error}$$

$$a^*(t) = |V(f;t) - V(o;t-1)| \quad \text{for } v \text{ - wind component error}$$

$$r(t) = \left[ (U(o;t) - U(o;t-1))^2 + (V(o;t) - V(o;t-1))^2 \right]^{1/2}$$

$$w(o;t) = \left[ U(o;t)^2 + V(o;t)^2 \right]^{1/2}$$

$$w(f;t) = \left[ U(f;t)^2 + V(f;t)^2 \right]^{1/2}$$

$$m(f;t) = \max(U(f;t), V(f;t))$$

TABLE 2B

Log-Likelihood  
500 mb. Height  
February Data

## Two-Variate Models

Wind Comp	Data	Model	Const	$r^*(t), a(o;t)$	$r(f;t), a(f;t)$	$r^*(t), a^*(t)$	$r^*(t), a(f;t)$	$r(t), w(o;t)$	$r(t), w(f;t)$
u	A	A	-7894.4	-7693.4	-7873.2	-7860.2	-7860.3	-7718.3	-7724.1
	B	B	-7965.1	-7676.8	-7923.7	-7807.4	-7813.6	-7762.7	-7772.2
	B	A	-7966.4	-7689.2	-7950.9	-7843.4	-7873.3	-7763.1	-7772.6
	A	B	-7895.6	-7704.0	-7893.6	-7891.5	-7905.7	-7718.7	-7724.5
v	A	A	-7720.5	-7622.3	-7706.6	-7695.7	-7701.5	-7601.9	-7609.0
	B	B	-7849.7	-7722.0	-7845.0	-7841.3	-7841.9	-7711.0	-7730.8
	B	A	-7853.8	-7733.8	-7853.9	-7853.6	-7850.4	-7716.5	-7734.6
	A	B	-7724.5	-7633.9	-7714.1	-7706.2	-7708.9	-7607.0	-7612.6

$$r(f;t) = \left[ (U(f;t) - U(f;t-1))^2 + (V(f;t) - V(f;t-1))^2 \right]^{1/2}$$

$$a(f;t) = |U(f;t) - U(f;t-1)|$$

$$r^*(t) = \left[ (U(f;t) - U(o;t-1))^2 + (V(f;t) - V(o;t-1))^2 \right]$$

$$a(o;t) = |U(o;t) - U(o;t-1)| \quad \text{for } u \text{ - wind component error}$$

$$a(o;t) = |V(o;t) - V(o;t-1)| \quad \text{for } v \text{ - wind component error}$$

$$a^*(t) = |U(f;t) - U(o;t-1)| \quad \text{for } u \text{ - wind component error}$$

$$a^*(t) = |V(f;t) - V(o;t-1)| \quad \text{for } v \text{ - wind component error}$$

$$r(t) = \left[ (U(o;t) - U(o;t-1))^2 + (V(o;t) - V(o;t-1))^2 \right]^{1/2}$$

$$w(o;t) = \left[ U(o;t)^2 + V(o;t)^2 \right]^{1/2}$$

$$w(f;t) = \left[ U(f;t)^2 + V(f;t)^2 \right]^{1/2}$$

TABLE 3B  
 One-Variate Models  
 Parameter Estimates  
 (Standard Errors)  
 500 mb. February Data

Wind Comp	Data Set	$r^*(t)$		$r(f;t)$		$a(o;t)$		$a(f;t)$		$a^*(t)$		$m(f;t)$	
		$\alpha$	$\beta$	$\alpha$	$\beta$	$\alpha$	$\beta$	$\alpha$	$\beta$	$\alpha$	$\beta$	$\alpha$	$\beta$
$u$	A	2.56 (0.06)	0.03 (0.005)	2.63 (0.06)	0.02 (0.006)	2.31 (0.05)	0.07 (0.006)	2.68 (0.05)	0.03 (0.008)	2.64 (0.05)	0.03 (0.007)	2.55 (0.06)	0.02 (0.003)
	B	2.25 (0.06)	0.06 (0.005)	2.60 (0.06)	0.02 (0.006)	2.28 (0.05)	0.08 (0.006)	2.83 (0.05)	0.01 (0.008)	2.50 (0.05)	0.06 (0.006)	2.53 (0.06)	0.02 (0.003)
	All	2.40 (0.04)	0.04 (0.004)	2.60 (0.04)	0.03 (0.004)	2.29 (0.03)	0.08 (0.004)	2.75 (0.03)	0.02 (0.006)	2.56 (0.03)	0.05 (0.005)	2.54 (0.04)	0.02 (0.002)
$v$	A	2.54 (0.06)	0.02 (0.005)	2.66 (0.06)	0.01 (0.006)	2.41 (0.05)	0.04 (0.005)	2.63 (0.05)	0.02 (0.007)	2.56 (0.05)	0.03 (0.006)	2.57 (0.06)	0.01 (0.003)
	B	2.69 (0.06)	0.01 (0.005)	2.75 (0.06)	0.008 (0.006)	2.40 (0.05)	0.05 (0.005)	2.75 (0.05)	0.01 (0.006)	2.72 (0.05)	0.01 (0.005)	2.58 (0.06)	0.01 (0.003)
	All	2.62 (0.04)	0.02 (0.004)	2.71 (0.04)	0.01 (0.004)	2.40 (0.03)	0.05 (0.003)	2.69 (0.03)	0.02 (0.004)	2.65 (0.03)	0.02 (0.004)	2.57 (0.04)	0.01 (0.002)

$$r(f;t) = \left[ (U(f;t) - U(f;t-1))^2 + (V(f;t) - V(f;t-1))^2 \right]^{1/2}$$

$$a(f;t) = |U(f;t) - U(f;t-1)|$$

$$r^*(t) = \left[ (U(f;t) - U(o;t-1))^2 + (V(f;t) - V(o;t-1))^2 \right]$$

$$a(o;t) = |U(o;t) - U(o;t-1)|, \quad a^*(t) = |U(f;t) - U(o;t-1)|$$



TABLE 4B

Two-Variate Models  
Parameter Estimates  
(Standard Errors)  
500 mb. February Data

Wind Comp	Data Set	$\log MSE = \alpha + \beta_1 r^*(t) + \beta_2 a(o;t)$		$\log MSE = \alpha + \beta_1 r_f(t) + \beta_2 a(f;t)$		$\log MSE = \alpha + \beta_1 r^*(t) + \beta_2 a^*(t)$				
		$\alpha$	$\beta_2$	$\alpha$	$\beta_2$	$\alpha$	$\beta_2$			
<i>u</i>	A	2.20 (0.07)	0.01 (0.005)	0.07 (0.006)	2.62 (0.06)	0.01 (0.008)	0.02 (0.01)	2.55 (0.06)	0.02 (0.007)	0.02 (0.009)
	B	2.02 (0.06)	0.03 (0.005)	0.07 (0.006)	2.58 (0.06)	0.05 (0.008)	-0.02 (0.01)	2.24 (0.06)	0.04 (0.007)	0.02 (0.008)
	All	2.10 (0.05)	0.02 (0.004)	0.07 (0.004)	2.60 (0.04)	0.03 (0.005)	-0.002 (0.007)	2.39 (0.04)	0.03 (0.005)	-0.02 (0.006)
<i>v</i>	A	2.33 (0.06)	0.009 (0.006)	0.04 (0.005)	2.68 (0.06)	-0.02 (0.01)	0.04 (0.01)	2.55 (0.06)	0.004 (0.008)	0.02 (0.009)
	B	2.45 (0.06)	-0.007 (0.006)	0.05 (0.005)	2.78 (0.06)	-0.009 (0.01)	0.02 (0.01)	2.69 (0.06)	0.008 (0.009)	0.007 (0.009)
	All	2.39 (0.05)	0.001 (0.004)	0.05 (0.004)	2.73 (0.04)	-0.01 (0.008)	0.03 (0.008)	2.62 (0.04)	0.006 (0.006)	0.01 (0.006)

$$r(f;t) = \left[ (U(f;t) - U(f;t-1))^2 + (V(f;t) - V(f;t-1))^2 \right]^{1/2}$$

$$a(f;t) = |U(f;t) - U(f;t-1)|$$

$$r^*(t) = \left[ (U(f;t) - U(o;t-1))^2 + (V(f;t) - V(o;t-1))^2 \right]$$

$$a(o;t) = |U(o;t) - U(o;t-1)|, \quad a^*(t) = |U(f;t) - U(o;t-1)|$$

TABLE 1C  
Log-Likelihood  
250 mb. Height/February Data

One-Variate Models

Wind Comp	Data Model	Const	$r^*(t)$	$r(f;t)$	$a(o;t)$	$a(f;t)$	$a^*(t)$	$r(t)$	$w(o;t)$	$w(f;t)$	$m(f;t)$	$U(f;t)$ $V(f;t)$	
u	A	-9376.4	-9253.8	-9365.4	-8609.8	-9362.6	-9289.9	-8705.7	-8614.6	-9137.6	-8962.3	-9074.1	
	B	-8957.8	-8897.4	-8927.6	-8559.6	-8950.7	-8946.3	-8637.6	-8600.3	-8890.9	-8795.1	-8911.9	
	B	A	-9001.3	-8931.4	-8977.3	-8562.1	-9024.6	-8989.3	-8641.4	-8611.9	-8923.2	-8816.7	-8991.9
	A	B	-9426.8	-9292.8	-9423.7	-8612.3	-9463.7	-9351.3	-8710.3	-8629.4	-9188.4	-8994.9	-9186.6
v	A	-8803.2	-8669.4	-8801.6	-8477.6	-8782.9	-8621.7	-8435.0	-8355.6	-8692.4	-8693.1	-8786.2	
	B	-8916.9	-8835.6	-8864.4	-8566.5	-8833.5	-8829.6	-8565.4	-8507.1	-8886.9	-8877.3	-8913.0	
	B	A	-8920.4	-8843.7	-8898.1	-8575.9	-8843.0	-8860.6	-8571.7	-8516.4	-8907.8	-8890.4	-8922.6
	A	B	-8806.6	-8677.3	-8820.1	-8486.7	-8788.3	-8647.8	-8440.6	-8364.2	-8714.4	-8706.5	-8794.8

$$r(f;t) = \left[ (U(f;t) - U(f;t-1))^2 + (V(f;t) - V(f;t-1))^2 \right]^{1/2}$$

$$a(f;t) = |U(f;t) - U(f;t-1)|$$

$$r^*(t) = \left[ (U(f;t) - U(o;t-1))^2 + (V(f;t) - V(o;t-1))^2 \right]$$

$$a(o;t) = |U(o;t) - U(o;t-1)|$$

$$a(o;t) = |V(o;t) - V(o;t-1)|$$

$$a^*(t) = |U(f;t) - U(o;t-1)|$$

$$a^*(t) = |V(f;t) - V(o;t-1)|$$

for u - wind component error

for v - wind component error

for u - wind component error

for v - wind component error

$$r(t) = \left[ (U(o;t) - U(o;t-1))^2 + (V(o;t) - V(o;t-1))^2 \right]^{1/2}$$

$$w(o;t) = \left[ U(o;t)^2 + V(o;t)^2 \right]^{1/2}$$

$$w(f;t) = \left[ U(f;t)^2 + V(f;t)^2 \right]^{1/2}$$

$$m(f;t) = \max(U(f;t), V(f;t))$$

TABLE 2C

Log-Likelihood  
250 mb. Height  
February Data

## Two-Variate Models

Wind Comp	Data Model	Const	$r^*(t), a(o;t)$	$r(f;t), a(f;t)$	$r^*(t), a^*(t)$	$r^*(t), a(f;t)$	$r(t), w(o;t)$	$r(t), w(f;t)$
u	A	-9376.4	-8537.9	-9361.5	-9252.4	-9231.0	-8504.0	-8685.4
	B	-8957.8	-8553.5	-8900.9	-8896.2	-8864.2	-8502.1	-8581.6
	B	-9001.3	-8572.7	-9004.9	-8934.7	-8981.0	-8507.2	-8594.4
	A	-9426.8	-8567.0	-9479.4	-9297.6	-9465.5	-8509.3	-8698.6
v	A	-8803.2	-8417.8	-8756.1	-8621.6	-8660.9	-8288.9	-8426.5
	B	-8916.9	-8565.9	-8828.3	-8821.9	-8818.3	-8415.8	-8541.6
	B	-8920.4	-8598.2	-8854.5	-8858.3	-8894.1	-8424.4	-8553.3
	A	-8806.6	-8469.7	-8772.4	-8646.8	-8710.8	-8296.5	-8437.9

$$r(f;t) = \left[ (U(f;t) - U(f;t-1))^2 + (V(f;t) - V(f;t-1))^2 \right]^{1/2}$$

$$a(f;t) = |U(f;t) - U(f;t-1)|$$

$$r^*(t) = \left[ (U(f;t) - U(o;t-1))^2 + (V(f;t) - V(o;t-1))^2 \right]$$

$$a(o;t) = |U(o;t) - U(o;t-1)| \quad \text{for } u\text{-wind component error}$$

$$a(o;t) = |V(o;t) - V(o;t-1)| \quad \text{for } v\text{-wind component error}$$

$$a^*(t) = |U(f;t) - U(o;t-1)| \quad \text{for } u\text{-wind component error}$$

$$a^*(t) = |V(f;t) - V(o;t-1)| \quad \text{for } v\text{-wind component error}$$

$$r(t) = \left[ (U(o;t) - U(o;t-1))^2 + (V(o;t) - V(o;t-1))^2 \right]^{1/2}$$

$$w(o;t) = \left[ U(o;t)^2 + V(o;t)^2 \right]^{1/2}$$

$$w(f;t) = \left[ U(f;t)^2 + V(f;t)^2 \right]^{1/2}$$

TABLE 3C  
One-Variate Models  
Parameter Estimates  
(Standard Errors)  
250 mb. February Data

Wind Comp	Data Set	$r^*(t)$		$r(f;t)$		$a(o;t)$		$a(f;t)$		$a^*(t)$		$m(f;t)$	
		$\alpha$	$\beta$	$\alpha$	$\beta$	$\alpha$	$\beta$	$\alpha$	$\beta$	$\alpha$	$\beta$	$\alpha$	$\beta$
u	A	3.50	0.03	3.81	0.02	3.10	0.06	3.88	0.02	3.72	0.03	3.00	0.03
		(0.06)	(0.003)	(0.07)	(0.005)	(0.05)	(0.004)	(0.05)	(0.005)	(0.04)	(0.003)	(0.06)	(0.002)
	B	3.34	0.03	3.52	0.02	3.07	0.05	3.88	-0.01	3.66	0.01	3.16	0.02
		(0.06)	(0.004)	(0.06)	(0.004)	(0.05)	(0.004)	(0.05)	(0.005)	(0.05)	(0.004)	(0.07)	(0.002)
	All	3.42	0.03	3.67	0.02	3.08	0.05	3.87	0.005	3.68	0.02	3.05	0.03
		(0.04)	(0.002)	(0.04)	(0.003)	(0.03)	(0.003)	(0.03)	(0.004)	(0.03)	(0.003)	(0.05)	(0.001)
v	A	3.07	0.04	3.63	0.005	3.13	0.04	3.50	0.02	3.10	0.05	3.14	0.02
		(0.06)	(0.005)	(0.06)	(0.004)	(0.04)	(0.003)	(0.05)	(0.005)	(0.05)	(0.004)	(0.06)	(0.002)
	B	3.22	0.03	3.42	0.02	3.10	0.04	3.44	0.03	3.37	0.03	3.34	0.01
		(0.07)	(0.004)	(0.06)	(0.004)	(0.05)	(0.002)	(0.05)	(0.004)	(0.05)	(0.004)	(0.07)	(0.002)
	All	3.15	0.03	3.46	0.03	3.12	0.04	3.46	0.03	3.25	0.04	3.23	0.02
		(0.05)	(0.003)	(0.03)	(0.003)	(0.03)	(0.002)	(0.03)	(0.003)	(0.04)	(0.003)	(0.05)	(0.001)

$$r(f;t) = \left[ (U(f;t) - U(f;t-1))^2 + (V(f;t) - V(f;t-1))^2 \right]^{1/2}$$

$$a(f;t) = |U(f;t) - U(f;t-1)|$$

$$r^*(t) = \left[ (U(f;t) - U(o;t-1))^2 + (V(f;t) - V(o;t-1))^2 \right]$$

$$a(o;t) = |U(o;t) - U(o;t-1)|, \quad a^*(t) = |U(f;t) - U(o;t-1)|$$

$$m(f;t) = \max(U(f;t), V(f;t))$$

TABLE 4C

Two-Variate Models  
 Parameter Estimates  
 (Standard Errors)  
 250 mb. February Data

Wind Comp	Data Set	log MSE = $\alpha + \beta_1 r^*(t) + \beta_2 a(o;t)$		log MSE = $\alpha + \beta_1 r_f(t) + \beta_2 a(f;t)$		log MSE = $\alpha + \beta_1 r^*(t) + \beta_2 a^*(t)$				
		$\alpha$	$\beta_2$	$\alpha$	$\beta_2$	$\alpha$	$\beta_2$			
u	A	2.73 (0.06)	0.03 (0.003)	0.05 (0.003)	3.83 (0.07)	0.007 (0.006)	0.01 (0.007)	3.50 (0.06)	0.03 (0.004)	0.006 (0.005)
	B	2.93 (0.07)	0.009 (0.004)	0.05 (0.004)	3.61 (0.06)	0.03 (0.004)	-0.03 (0.006)	3.35 (0.06)	0.03 (0.004)	-0.006 (0.005)
	All	2.80 (0.05)	0.02 (0.002)	0.05 (0.003)	3.68 (0.04)	0.02 (0.004)	-0.01 (0.004)	3.42 (0.04)	0.03 (0.003)	0.007 (0.004)
v	A	2.80 (0.06)	0.02 (0.003)	0.03 (0.003)	3.69 (0.07)	-0.04 (0.008)	0.06 (0.008)	3.09 (0.06)	0.001 (0.005)	0.05 (0.006)
	B	3.06 (0.07)	0.003 (0.004)	0.04 (0.003)	3.52 (0.06)	-0.02 (0.007)	0.05 (0.007)	3.25 (0.07)	0.02 (0.006)	0.02 (0.006)
	All	2.90 (0.04)	0.02 (0.002)	0.04 (0.002)	3.58 (0.04)	-0.03 (0.005)	0.05 (0.005)	3.18 (0.04)	0.01 (0.004)	0.03 (0.004)

$$r(f;t) = \left[ (u(f;t) - u(f;t-1))^2 + (v(f;t) - v(f;t-1))^2 \right]^{1/2}$$

$$a(f;t) = |u(f;t) - u(f;t-1)|$$

$$r^*(t) = \left[ (u(f;t) - u(o;t-1))^2 + (v(f;t) - v(o;t-1))^2 \right]$$

$$a(o;t) = |u(o;t) - u(o;t-1)|, \quad a^*(t) = |u(f;t) - u(o;t-1)|$$

TABLE 5  
Percent of Increase  
 $(\bar{\ell} - \bar{\ell}_c) / |\bar{\ell}_c|$

One-Variate Models

Pressure Level mb	Wind Comp	Data Set	Model	$r$	$w(o)$	$a(o)$	$r^*$	$a^*$	$w(f)$	$m(f)$
850	$u$	B	A	0.8	1.6	1.5	1.4	1.9	0.2	-0.2
			B	0.8	0.5	1.7	0.5	0.8	-0.03	-0.1
	$v$	A	A	0.7	2.2	0.4	0.7	0.3	0.2	0.1
			B	0.4	2.6	-0.3	1.3	1.1	0.08	-0.09
500	$u$	B	A	2.3	0.9	3.2	1.4	1.2	0.3	0.03
			B	2.0	0.6	2.4	-0.04	0.2	0.3	0.4
	$v$	A	A	1.3	0.8	1.6	0.006	-0.005	0.2	0.2
			B	1.3	0.4	1.2	0.1	0.2	0.3	0.1
250	$u$	B	A	3.5	3.9	4.4	0.3	-0.4	0.4	3.0
			B	7.1	8.0	8.1	0.8	0.3	2.0	4.0
	$v$	A	A	3.9	4.5	3.8	0.8	0.6	0.1	0.3
			B	4.1	5.0	3.6	1.4	1.7	1.0	0.1

**TABLE 6**  
**Sample Moments of**  
**First-Guess Wind Speed Errors Divided by Predicted Standard Deviations**

Wind Comp	Covariates	Est. Method	Mean	Std.Dev.	Skewness	Kurtosis	
u	none	-	0.11	7.04	-1.28	51.1	
	constant	MLE	0.02	1.01	-1.24	57.4	
	w(f)	MLE	0.02	1.01	-0.73	32.9	
	w(f)	2 bins	0.02	1.01	-0.59	34.4	
		3 bins	0.02	1.02	-0.98	42.9	
	m(f)	MLE	0.01	1.01	1.01	27.6	
	m(f)	2 bins	0.02	1.01	-0.63	34.1	
		3 bins	0.01	1.03	-1.11	41.0	
	a(o)	MLE	0.02	1.00	0.24	6.0	
	r(o)	MLE	0.02	1.00	0.17	7.7	
	a(f)	MLE	0.02	1.02	-0.92	57.6	
	a*	MLE	0.01	1.00	-1.42	64.5	
	v	none	-	-0.11	6.47	-1.98	35.5
		constant	MLE	-0.02	1.00	-1.98	35.3
w(f)		MLE	-0.00	1.01	-1.49	28.6	
w(f)		2 bins	-0.00	1.00	-1.28	23.9	
		3 bins	-0.01	1.03	-2.04	37.8	
m(f)		MLE	-0.00	1.00	-1.33	24.8	
m(f)		2 bins	-0.02	1.00	-2.00	36.2	
		3 bins	-0.01	1.04	-2.14	40.9	
a(o)		MLE	-0.02	1.00	2.08	36.1	

$$m(f;t) = \max(U(f;t), V(f;t))$$

$$w(f;t) = [U(f;t)^2 + V(f;t)^2]^{1/2}$$

## INITIAL DISTRIBUTION LIST

1. Research Office (Code 08) .....1  
Naval Postgraduate School  
Monterey, CA 93943-5000
2. Dudley Knox Library (Code 52) .....2  
Naval Postgraduate School  
Monterey, CA 93943-5002
3. Defense Technical Information Center.....2  
Cameron Station  
Alexandria, VA 22314
4. Department of Operations Research (Code OR).....1  
Naval Postgraduate School  
Monterey, CA 93943-5000
5. Prof. Donald Gaver (Code OR/Gv) .....5  
Naval Postgraduate School  
Monterey, CA 93943-5000
6. Prof. Patricia Jacobs (Code OR/Jc) .....5  
Naval Postgraduate School  
Monterey, CA 93943-5000
7. Center for Naval Analyses.....1  
4401 Ford Avenue  
Alexandria, VA 22302-0268
8. Dr. J. Abrahams, Code 1111, Room 607 .....1  
Mathematical Sciences Division, Office of Naval Research  
800 North Quincy Street  
Arlington, VA 22217-5000
9. Dr. David Brillinger.....1  
Statistics Department  
University of California  
Berkeley, CA 94720



10. Prof. Brad Carlin.....1  
 School of Public Health  
 University of Minnesota  
 Mayo Bldg. A460  
 Minneapolis, MN 55455
11. Prof. H. Chernoff.....1  
 Department of Statistics  
 Harvard University  
 1 Oxford Street  
 Cambridge, MA 02138
12. Dr. John Copas.....1  
 Dept. of Mathematics,  
 University of Birmingham  
 P. O. Box 363  
 Birmingham B15 2TT  
 ENGLAND
13. Professor Sir David Cox.....1  
 Nuffield College  
 Oxford, OX1 1NF  
 ENGLAND
14. Professor H. G. Daellenbach.....1  
 Department of Operations Research  
 University of Canterbury  
 Christchurch, NEW ZEALAND
15. Dr. S. R. Dalal.....1  
 Bellcore  
 445 South Street  
 Morristown, NJ 07962-1910
16. Dr. D. F. Daley .....1  
 Statistic Dept. (I.A.S.)  
 Australian National University  
 Canberra, A.C.T. 2606  
 AUSTRALIA

17.	Prof. Bradley Efron.....	1
	Statistics Dept. Sequoia Hall Stanford University Stanford, CA 94305	
18.	Prof. George S. Fishman.....	1
	Curr. in OR & Systems Analysis University of North Carolina Chapel Hill, NC 20742	
19.	Dr. Neil Gerr .....	1
	Office of Naval Research Arlington, VA 22217	
20.	Dr. J. Goerss .....	10
	Naval Oceanographic and Atmospheric Laboratory Monterey, CA 93943-5006	
21.	Dr. D. C. Hoaglin .....	1
	Department of Statistics Harvard University 1 Oxford Street Cambridge, MA 02138	
22.	Institute for Defense Analysis.....	1
	1800 North Beauregard Alexandria, VA 22311	
23.	Prof. J. B. Kadane.....	1
	Dept. of Statistics Carnegie-Mellon University Pittsburgh, PA 15213	
24.	Dr. Jon Kettenring.....	1
	Bellcore 445 South Street Morris Township, NJ 07962-1910	
25.	Koh Peng Kong.....	1
	OA Branch, DSO Ministry of Defense Blk 29 Middlesex Road SINGAPORE 1024	

26. Dr. A. J. Lawrence.....1  
 Dept. of Mathematics,  
 University of Birmingham  
 P. O. Box 363  
 Birmingham B15 2TT  
 ENGLAND
27. Prof. M. Leadbetter.....1  
 Department of Statistics  
 University of North Carolina  
 Chapel Hill, NC 27514
28. Prof. J. Lehoczky.....1  
 Department of Statistics  
 Carnegie-Mellon University  
 Pittsburgh, PA 15213
29. Dr. Colin Mallows.....1  
 AT&T Bell Telephone Laboratories  
 600 Mountain Avenue  
 Murray Hill, NJ 07974
30. Prof. M. Mazumdar.....1  
 Dept. of Industrial Engineering  
 University of Pittsburgh  
 Pittsburgh, PA 15235
31. Prof. Carl N. Morris.....1  
 Statistics Department  
 Harvard University  
 1 Oxford St.  
 Cambridge, MA 02138
32. Prof. F. W. Mosteller.....1  
 Department of Statistics  
 Harvard University  
 1 Oxford St.  
 Cambridge, MA 02138
33. Operations Research Center, Rm E40-164.....1  
 Massachusetts Institute of Technology  
 Attn: R. C. Larson and J. F. Shapiro  
 Cambridge, MA 02139

34. Prof. Frank Samaniego .....1  
 Statistics Department  
 University of California  
 Davis, CA 95616
35. Prof. W. R. Schucany.....1  
 Dept. of Statistics  
 Southern Methodist University  
 Dallas, TX 75222
36. Prof. H. Solomon .....1  
 Department of Statistics  
 Sequoia Hall  
 Stanford University  
 Stanford, CA 94305
37. Prof. W. Stuetzle .....1  
 Department of Statistics  
 University of Washington  
 Seattle, WA 98195
38. Prof. J. R. Thompson.....1  
 Dept. of Mathematical Science  
 Rice University  
 Houston, TX 77001
39. Prof. J. W. Tukey .....1  
 Statistics Dept., Fine Hall  
 Princeton University  
 Princeton, NJ 08540
40. Dr. D. Vere-Jones.....1  
 Dept. of Math, Victoria Univ. of Wellington  
 P. O. Box 196  
 Wellington  
 NEW ZEALAND
41. Prof. David L. Wallace .....1  
 Statistics Dept., University of Chicago  
 5734 S. University Ave.  
 Chicago, IL 60637

- 42. Dr. Ed Wegman.....1  
George Mason University  
Fairfax, VA 22030
- 43. Dr. P. Welch.....1  
IBM Research Laboratory  
Yorktown Heights, NY 10598
- 44. Prof. Roy Welsch.....1  
Sloan School  
M.I.T.  
Cambridge, MA 02139











DUDLEY KNOX LIBRARY



3 2768 00337188 1