

# Multiple Random Variables

Young W Lim

June 8, 2020

Copyright (c) 2018 Young W. Lim. Permission is granted to copy, distribute and/or modify this document under the terms of the GNU Free Documentation License, Version 1.2 or any later version published by the Free Software Foundation; with no Invariant Sections, no Front-Cover Texts, and no Back-Cover Texts. A copy of the license is included in the section entitled "GNU Free Documentation License".

This work is licensed under a Creative Commons "Attribution-NonCommercial-ShareAlike 3.0 Unported" license.



Based on  
Probability, Random Variables and Random Signal Principles,  
P.Z. Peebles, Jr. and B. Shi

# Outline

- 1 Central Limit Theorem
  - Unequal Distributions
  - Equal Distributions

# Outline

- 1 Central Limit Theorem
  - Unequal Distributions
  - Equal Distributions

# Central Limit Theorem

## Definition

the central limit theorem says that the probability distribution function of the sum of large number of random variables approaches a Gaussian distribution.

This theorem is known to apply some cases of statistically independent random variables.

# Central Limit Theorem

## Unequal Distribution Case

### Definition

the sum  $Y$  of  $N$  independent random variables  $X_1, X_2, \dots, X_N$   
Let  $Y = X_1 + X_2 + \dots + X_N$ , then

$$\bar{Y}_N = \bar{X}_1 + \bar{X}_2 + \dots + \bar{X}_N$$

$$\sigma_{Y_N}^2 = \sigma_{X_1}^2 + \sigma_{X_2}^2 + \dots + \sigma_{X_N}^2$$

the probability distribution of  $Y$  asymptotically approaches to Gaussian distribution function as  $N \rightarrow \infty$

# Sufficient Conditions

## Unequal Distribution Case

### Definition

$$\sigma_{X_i}^2 > B_1 > 0 \quad i = 1, 2, \dots, N$$

$$E[|X_i - \bar{X}_i|^3] < B_2 \quad i = 1, 2, \dots, N$$

where  $B_1$  and  $B_2$  are positive numbers

these conditions guarantee that no one random variable in the sum dominates

# Distribution vs density functions

## Unequal Distribution Case

the central limit theorem guarantees

- only that the distribution of the sum of random variables become Gaussian
- the density of the sum of random variables is not always Gaussian
- the sum of continuous random variables :
  - under certain conditions on individual random variables the density of the sum is always Gaussian
- the sum of discrete random variables :
  - the density function may contain impulses and thus is not Gaussian.



# Discrete Random Variable Examples

distribution may contain impulses

the sum  $Y$  of  $N$  independent discrete random variables

$$Y = X_1 + X_2 + \dots + X_N$$

- discrete random variable
- density function may contain impulses
- therefore the density function is not Gaussian
- although the distribution approaches Gaussian
  
- when the possible discrete values of each random variable are  $kb, k = 0, \pm 1, \pm 2, \dots$ , where  $b$  is a constant
  - the envelope of the impulses in the density of the sum will be Gaussian
  - with the mean  $Y_N$  and variance  $\sigma_{Y_N}^2$

# Central Limit Theorem (1)

## Equal Distribution Case

### Definition

the sum  $Y$  of  $N$  independent random variables  $X_1, X_2, \dots, X_N$   
assume that  $X_1, X_2, \dots, X_N$  have the same distribution function.

Let  $Y_N = X_1 + X_2 + \dots + X_N$ ,

then  $W_N = (Y_N - \bar{Y}_N) / \sigma_{Y_N}$  is

the zero-mean, unit-variance random variable

## Central Limit Theorem (2)

## Equal Distribution Case

## Definition

Let  $Y_N = X_1 + X_2 + \cdots + X_N$  and  $W_N = (Y_N - \bar{Y}_N) / \sigma_{Y_N}$

$$\begin{aligned} W_N &= (Y_N - \bar{Y}_N) / \sigma_{Y_N} \\ &= \sum_{i=1}^N (X_i - \bar{X}_i) / \left[ \sum_{i=1}^N \sigma_{\bar{X}_i}^2 \right]^{1/2} \\ &= \frac{1}{\sqrt{N} \sigma_X} \sum_{i=1}^N (X_i - \bar{X}_i) \end{aligned}$$

where  $X_i = \bar{X}$  and  $\sigma_{\bar{X}_i}^2 = \sigma_{\bar{X}}$

## Central Limit Theorem (3)

## Equal Distribution Case

$$\begin{aligned}
 W_N &= (Y_N - \bar{Y}_N) / \sigma_{Y_N} \\
 &= \left( X_i - \sum_{i=1}^N \bar{X}_i \right) / \left[ \sum_{i=1}^N \sigma_{\bar{X}_i}^2 \right]^{1/2} \\
 &= \sum_{i=1}^N (X_i - \bar{X}_i) / \left[ \sum_{i=1}^N \sigma_{\bar{X}_i}^2 \right]^{1/2} \\
 &= \sum_{i=1}^N (X_i - \bar{X}_i) / [N\sigma_X^2]^{1/2} \\
 &= \frac{1}{\sqrt{N}\sigma_X} \sum_{i=1}^N (X_i - \bar{X}_i)
 \end{aligned}$$

where  $\bar{X}_i = \bar{X}$  and  $\sigma_{\bar{X}_i}^2 = \sigma_{\bar{X}}^2$

$\bar{Y}_N = \bar{X}_1 + \bar{X}_2 + \dots + \bar{X}_N$  and  $\sigma_{Y_N}^2 = \sigma_{X_1}^2 + \sigma_{X_2}^2 + \dots + \sigma_{X_N}^2$

# Characteristic Function(1)

## Equal Distribution Case

the characteristic function of  $W_N$   
a zero mean, unit variance Gaussian random variable

$$\Phi_{W_N}(\omega) = \exp(-\omega^2/2)$$

$W_N$  is the density of the Gaussian random variable  
Fourier transforms are unique

$$\Phi_{W_N}(\omega) = E[e^{j\omega W_N}]$$

## Characteristic Function(2)

## Equal Distribution Case

$$W_N = \frac{1}{\sqrt{N}\sigma_X} \sum_{i=1}^N (X_i - \bar{X}_i)$$

$$\begin{aligned}\Phi_{W_N}(\omega) &= E[e^{j\omega W_N}] = E\left[\exp\left(\frac{j\omega}{\sqrt{N}\sigma_X} \sum_{i=1}^N (X_i - \bar{X})\right)\right] \\ &= E\left[\exp\left(\frac{j\omega}{\sqrt{N}\sigma_X} (X_1 - \bar{X})\right) \cdots \exp\left(\frac{j\omega}{\sqrt{N}\sigma_X} (X_N - \bar{X})\right)\right] \\ &= \left\{ E\left[\exp\left(\frac{j\omega}{\sqrt{N}\sigma_X} (X_1 - \bar{X})\right)\right] \right\}^N\end{aligned}$$

$$E[X_1] = E[X_2] = \cdots = E[X_N]$$

# Characteristic Function (3)

## Equal Distribution Case

$$\Phi_{W_N}(\omega) = \left\{ E \left[ \exp \left( \frac{j\omega}{\sqrt{N}\sigma_X} (X_1 - \bar{X}) \right) \right] \right\}^N$$
$$\ln[\Phi_{W_N}(\omega)] = N \ln \left\{ E \left[ \exp \left( \frac{j\omega}{\sqrt{N}\sigma_X} (X_1 - \bar{X}) \right) \right] \right\}$$

$$E \left[ \exp \left( \frac{j\omega}{\sqrt{N}\sigma_X} (X_1 - \bar{X}) \right) \right]$$
$$= E \left[ 1 + \frac{j\omega}{\sqrt{N}\sigma_X} (X_1 - \bar{X}) + \left( \frac{j\omega}{\sqrt{N}\sigma_X} \right)^2 (X_1 - \bar{X})^2 + \frac{R_N}{N} \right]$$
$$= 1 - \frac{\omega^2}{2N} + \frac{E[R_N]}{N}$$

where  $E[R_N]$  approaches zero as  $N \rightarrow \infty$

# Characteristic Function (4)

## Equal Distribution Case

$$\ln[\Phi_{W_N}(\omega)] = N \ln \left[ 1 - \frac{\omega^2}{2N} + \frac{E[R_N]}{N} \right]$$

$$\ln[1 - z] = - \left[ z + \frac{z^2}{2} + \frac{z^3}{3} + \dots \right], |z| < 1$$

$$\ln[\Phi_{W_N}(\omega)] = -\frac{\omega^2}{2} + E[R_N] - \frac{N}{2} \left[ \frac{\omega^2}{2N} + \frac{E[R_N]}{N} \right]^2 + \dots$$

$$\lim_{N \rightarrow \infty} \ln[\Phi_{W_N}(\omega)] = \ln \left[ \lim_{N \rightarrow \infty} \Phi_{W_N}(\omega) \right] = -\frac{\omega^2}{2}$$

$$\lim_{N \rightarrow \infty} \Phi_{W_N}(\omega) = e^{-\frac{\omega^2}{2}}$$