# Characteristics of Multiple Random Variables 

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Based on
Probability, Random Variables and Random Signal Principles, P.Z. Peebles,Jr. and B. Shi

## Outline

(1) Joint Guassian Random Variables

## Bivariate Gaussian Density

## Definition

The two random variables $X$ and $Y$ are said to be jointly Gaussian, if their joint density function is

$$
\begin{gathered}
f_{X, Y}(x, y)=\frac{1}{2 \pi \sigma_{X} \sigma_{Y} \sqrt{1-\rho^{2}}} \cdot \\
\exp \left\{\frac{-1}{2\left(1-\rho^{2}\right)} \cdot\left[\frac{(x-\bar{X})^{2}}{\sigma_{X}^{2}}-\frac{2 \rho(x-\bar{X})(y-\bar{Y})}{\sigma_{X} \sigma_{Y}}+\frac{(y-\bar{Y})^{2}}{\sigma_{Y}^{2}}\right]\right\} \\
\bar{X}=E[X], Y=E[Y], \sigma_{X}^{2}=E\left[(X-\bar{X})^{2}\right], \sigma_{Y}^{2}=E\left[(Y-\bar{Y})^{2}\right], \\
\rho=E[(X-\bar{X})(Y-\bar{Y})] / \sigma_{X} \sigma_{Y}
\end{gathered}
$$

## Bivariate Gaussian Density - Maximum value

$$
f_{X, Y}(x, y) \leq f_{X, Y}(\bar{X}, \bar{Y})=\frac{1}{2 \pi \sigma_{X} \sigma_{Y} \sqrt{1-\rho^{2}}}
$$

## Bivariate Gaussian Density - Uncorrelated

two random variables
$f_{X, Y}(x, y)=f_{X}(x) f_{Y}(x)$ is sufficient to guarantee that $X$ and $Y$ are statistically independent. Any uncorrelated Guassian random variables are also statistically independent a coordinate rotation (linear transformation of $X$ and $Y$ ) through the angle

$$
\theta=\frac{1}{2} \tan ^{-1}\left[\frac{2 \rho \sigma_{X} \sigma_{Y}}{\sigma_{X}^{2} \sigma_{Y}^{2}}\right]
$$

is sufficient to convert correlated random variables $X$ and $Y$ having $\sigma_{X}^{2}$ and $\sigma_{Y}^{2}$, respectively, correlation coefficient $\rho$, and the joint densityof $f_{X, Y}(x, y)=\frac{1}{2 \pi \sigma_{X} \sigma_{Y} \sqrt{1-\rho^{2}}} \cdot \exp [\cdots]$ into two statistically independent Gaussian random variables

## Multi-variate Gaussian Density

$N$ random variables $X_{1}, X_{2}, \ldots, X_{N}$ are called jointly Gaussian if their joint density function can be written as

$$
\begin{gathered}
f_{X_{1}, \cdots, x_{N}}\left(x_{1}, \cdots, x_{N}\right)=\frac{\mid\left[\left.\left.C_{X}\right|^{-1}\right|^{1 / 2}\right.}{(2 \pi)^{N / 2}} \exp \left\{-\frac{[x-\bar{X}]^{t}\left[C_{X}\right][x-\bar{X}]}{2}\right\} \\
{[x-\bar{X}]=\left[\begin{array}{c}
x_{1}-\bar{X}_{1} \\
x_{2}-\bar{X}_{2} \\
\\
x_{N}-\bar{X}_{N}
\end{array}\right], \quad\left[C_{X}\right]=\left[\begin{array}{cccc}
C_{11} & C_{12} & \cdots & C_{1 N} \\
C_{21} & C_{22} & \cdots & C_{2 N} \\
\vdots & \vdots & \ddots & \vdots \\
C_{N 1} & C_{N 2} & \cdots & C_{N N}
\end{array}\right]}
\end{gathered}
$$

## Multi-variate Gaussian Density - notations $N$ random variables

$N$ random variables $X_{1}, X_{2}, \ldots, X_{N}$ are called jointly Gaussian if their joint density function can be written as

$$
f_{X_{1}, \cdots, X_{N}}\left(x_{1}, \cdots, x_{N}\right)=\frac{\mid\left[\left.\left.C_{X}\right|^{-1}\right|^{1 / 2}\right.}{(2 \pi)^{N / 2}} \exp \left\{-\frac{[x-\bar{X}]^{t}\left[C_{X}\right][x-\bar{X}]}{2}\right\}
$$

where $[\bullet]^{t}$ denotes a matrix transposition,
$[\bullet]^{-1}$ denotes a matrix inversion


## Covariance Matrix

## $N$ random variables

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$$
f_{X_{1}, \cdots, X_{N}}\left(x_{1}, \cdots, x_{N}\right)=\frac{\mid\left[\left.\left.C_{X}\right|^{-1}\right|^{1 / 2}\right.}{(2 \pi)^{N / 2}} \exp \left\{-\frac{[x-\bar{X}]^{t}\left[C_{X}\right][x-\bar{X}]}{2}\right\}
$$

where $\left[C_{x}\right]$ is called the covariance matrix of $N$ random variables

$$
C_{i j}=E\left[\left(X_{i}-\bar{X}_{i}\right)\left(X_{j}-\bar{X}_{j}\right)\right]=\left\{\begin{array}{cl}
\sigma_{X_{i}}^{2} & i=j \\
C_{X_{i} X_{j}} & i \neq j
\end{array}\right.
$$

## Covariance Matrix ( $N=2$ )

## $N$ random variables

$$
\begin{gathered}
f_{X_{1} X_{2}}\left(x_{1}, x_{2}\right)=\frac{\mid\left[\left.\left.C_{X}\right|^{-1}\right|^{1 / 2}\right.}{(2 \pi)^{2 / 2}} \exp \left\{-\frac{[x-\bar{X}]^{t}\left[C_{X}\right][x-\bar{X}]}{2}\right\} \\
{\left[C_{X}\right]=\left[\begin{array}{cc}
C_{11} & C_{12} \\
C_{21} & C_{22}
\end{array}\right]=\left[\begin{array}{cc}
\sigma_{X_{1}}^{2} & \rho \sigma_{X_{1}} \sigma_{X_{2}} \\
\rho \sigma_{X_{1}} \sigma_{X_{2}} & \sigma_{X_{2}}^{2}
\end{array}\right]} \\
{\left[C_{X}\right]^{-1}=\frac{1}{1-\rho^{2}}\left[\begin{array}{cc}
\sigma_{X_{1}}^{2} & -\rho / \sigma_{X_{1}} \sigma_{X_{2}} \\
-\rho / \sigma_{X_{1}} \sigma_{X_{2}} & \sigma_{X_{2}}^{2}
\end{array}\right]} \\
\left|\left[C_{X}\right]^{-1}\right|=1 / \sigma_{X_{1}}^{2} \sigma_{X_{2}}^{2}\left(1-\rho^{2}\right)
\end{gathered}
$$

